

```

-----
-----
name: <unnamed>
log: C:\Users\nw\Documents\HILDA Project\Training\Melbourne Nov
2011\Exercises\ex6.log
log type: text
opened on: 8 Nov 2011, 09:37:21

```

```

.
. *****
. * Load the panel data and tell Stata it is panel data *
. * then, only keep males to reduce the sample size and increase*
. * the speed of estimation.
*
. *****
. use "`working'\panelex5.dta",replace

. tsset id wave
. panel variable: id (unbalanced)
. time variable: wave, 1 to 9, but with gaps
. delta: 1 unit

```

```

.
.
. *****
. * Continuing with the wage example, suppose we specifically *
. * want to examine the determinants of low pay. Create a dummy*
. * variable to indicate an hourly wage of less than $14/hour. *
. * Examine the proportion of low paid workers, how it has *
. * changed over time and low pay transitions.
. *****
.
. gen lopay=w_hr<14

```

```

. xttab lopay

```

lopay	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	47428	79.43	11431	85.55	85.06
1	12279	20.57	6420	48.05	56.68
Total	59707	100.00	17851	133.60	74.85

(n = 13362)

```

. xttrans lopay

```

lopay	lopay		Total
	0	1	
0	93.24	6.76	100.00
1	50.59	49.41	100.00
Total	84.30	15.70	100.00

```

. bysort wave: su lopay

```


-> wave = 1

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6686	.3305414	.4704433	0	1

-> wave = 2

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6382	.2919148	.4546789	0	1

-> wave = 3

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6371	.2605556	.4389723	0	1

-> wave = 4

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6258	.2245126	.4172943	0	1

-> wave = 5

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6602	.1990306	.3993014	0	1

-> wave = 6

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6787	.1753352	.3802816	0	1

-> wave = 7

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6816	.149061	.3561748	0	1

-> wave = 8

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6826	.1273074	.3333414	0	1

-> wave = 9

Variable	Obs	Mean	Std. Dev.	Min	Max
lopay	6979	.1077518	.3100889	0	1

```

.
. *****
. * Estimate a linear probability model (LPM) to explain low *
. * pay, using the same specification as the wage equation. *
. * Estimate a RE model. How would you interpret the *
. * coefficients? [Note they can be interpreted directly, *
. * unlike in the probit and logit models below.] *
. *****
. xtreg lopay hgage agesq cohort married female further degree ///
> jbemp /*tucov*/ permanent nsw, re

```

```

Random-effects GLS regression           Number of obs   =
59707
Group variable: id                     Number of groups =
13362

R-sq:  within = 0.1358                  Obs per group: min =
1                                           avg =
4.5                                       max =
9                                           overall = 0.2308

                                           Wald chi2(10)   =
13153.36
corr(u_i, X) = 0 (assumed)              Prob > chi2     =
0.0000

```

lopay	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-------	-------	-----------	---	------	----------------------

95%	1	1	Skewness	1.456512
99%	1	1	Kurtosis	3.121427

Linear prediction

Percentiles		Smallest		
1%	-.1330489	-.2241488		
5%	-.0695236	-.2169429		
10%	-.0247208	-.211746	Obs	59707
25%	.0621021	-.2044594	Sum of Wgt.	59707
			Mean	.2153759
50%	.1768914		Std. Dev.	.2102552
			Largest	
75%	.3292528	.8914966	Variance	.0442072
90%	.5359221	.8915645	Skewness	.7376855
95%	.6372729	.8917341	Kurtosis	3.054582
99%	.770086	.8918359		

Xb + u[id]

Percentiles		Smallest		
1%	-.1494125	-.2884432		
5%	-.0887211	-.2728885		
10%	-.0492079	-.2660058	Obs	59707
25%	.0224265	-.2659074	Sum of Wgt.	59707
			Mean	.2056543
50%	.1277282		Std. Dev.	.2517365
			Largest	
75%	.3411561	1.125741	Variance	.0633712
90%	.6035851	1.138452	Skewness	1.009622
95%	.7397249	1.150896	Kurtosis	3.306962
99%	.8936468	1.156993		

```

.
. *****
. * Estimate an FE LPM and examine the predicted probabilities.*
. *****
. xtreg lopay hgage agesq cohort married female further degree ///
>      jbemp /*tucov*/ permanent nsw, fe
note: cohort omitted because of collinearity
note: female omitted because of collinearity

```

Fixed-effects (within) regression	Number of obs	=	59707
Group variable: id	Number of groups	=	13362
R-sq: within = 0.1406	Obs per group: min	=	1
	avg	=	4.5
	max	=	9
	F(8,46337)	=	947.51
corr(u_i, Xb) = -0.8384	Prob > F	=	0.0000

99% 1 1 Kurtosis 3.121427

 Linear prediction

	Percentiles	Smallest		
1%	-.4637284	-.4865073		
5%	-.4218718	-.4864751		
10%	-.3901945	-.4863298	Obs	59707
25%	-.2765585	-.4861845	Sum of Wgt.	59707
50%	.0331033		Mean	.2056543
		Largest	Std. Dev.	.5589684
75%	.6036463	1.518717		
90%	1.098327	1.519557	Variance	.3124456
95%	1.261273	1.520396	Skewness	.7271015
99%	1.51344	1.520396	Kurtosis	2.329063

 Xb + u[id]

	Percentiles	Smallest		
1%	-.1859578	-.3880788		
5%	-.1072756	-.3713734		
10%	-.063895	-.3487845	Obs	59707
25%	-.005028	-.3476784	Sum of Wgt.	59707
50%	.0903401		Mean	.2056543
		Largest	Std. Dev.	.3051392
75%	.3457097	1.285256		
90%	.6799397	1.292453	Variance	.0931099
95%	.9556987	1.311241	Skewness	1.255784
99%	1.042445	1.408116	Kurtosis	3.79189

```

.
. *****
. * Estimate the model as a RE probit and examine the
. * predicted probabilities (type help xtprobit post-estimation*
. * for details of the predict command).
. * What do you notice? How have the probabilities been
. * calculated?
. *****
. xtprobit lopay hgage agesq cohort married female further degree ///
> jbemp /*tucov*/ permanent nsw, re
  
```

Fitting comparison model:

```

Iteration 0: log likelihood = -30339.617
Iteration 1: log likelihood = -23420.818
Iteration 2: log likelihood = -23251.665
Iteration 3: log likelihood = -23251.458
Iteration 4: log likelihood = -23251.458
  
```

Fitting full model:

```

rho = 0.0 log likelihood = -23251.458
rho = 0.1 log likelihood = -22388.762
rho = 0.2 log likelihood = -21989.801
rho = 0.3 log likelihood = -21836.779
rho = 0.4 log likelihood = -21852.112
  
```



```

      sigma_u | .9415761 .0173648 .9081496
.976233
      rho | .4699363 .0091878 .4519754
.4879753
-----
-----

```

Likelihood-ratio test of rho=0: $\chi^2(01) = 3767.43$ Prob $\geq \chi^2 = 0.000$

```

.
. predict reprob, pu0
. su lopay reprob, de

```

lopay

Percentiles		Smallest		
1%	0	0		
5%	0	0		
10%	0	0	Obs	59707
25%	0	0	Sum of Wgt.	59707
50%	0		Mean	.2056543
		Largest	Std. Dev.	.4041823
75%	0	1		
90%	1	1	Variance	.1633633
95%	1	1	Skewness	1.456512
99%	1	1	Kurtosis	3.121427

Pr(lopay=1 assuming u_i=0)

Percentiles		Smallest		
1%	.0002255	.0000104		
5%	.0012381	.0000139		
10%	.003309	.0000154	Obs	59707
25%	.0150408	.0000181	Sum of Wgt.	59707
50%	.0678746		Mean	.1749603
		Largest	Std. Dev.	.2321324
75%	.238481	.9754409		
90%	.5617821	.9755037	Variance	.0538855
95%	.7365823	.9756603	Skewness	1.662574
99%	.9089257	.9757539	Kurtosis	4.866044

```

.
. *****
. * Now estimate the model as a RE logit and examine the      *
. * predicted probabilities (type help xtlogit postestimation *
. * for details of the predict command).                      *
. * What do you notice? How have the probabilities been      *
. * calculated?                                              *
. *****
. xtlogit lopay hgage agesq cohort married female further degree ///
>      jbemp /*tucov*/ permanent nsw, re

```

Fitting comparison model:

Iteration 0: log likelihood = -30339.617


```

      degree | -1.300175   .0693114   -18.76   0.000   -1.436023   -
1.164327
      jbempt | -.0521376   .0038565   -13.52   0.000   -.0596961   -
.044579
      permanent | -.2926862   .0359879   -8.13   0.000   -.3632212   -
.2221512
      nsw | -.2433049   .0484902   -5.02   0.000   -.338344   -
.1482658
      _cons | 708.8378   13.90222   50.99   0.000   681.5899
736.0856

```

```

-----+-----
      /lnsig2u | 1.059181   .0375428   .9855987
1.132764

```

```

-----+-----
      sigma_u | 1.698237   .0318783   1.636892
1.761881
      rho | .4671309   .0093451   .4488672
.4854832

```

```

-----
Likelihood-ratio test of rho=0: chibar2(01) = 3734.11 Prob >= chibar2 = 0.000

```

```

.
. predict relprob, pu0
. su lopay relprob, de

```

lopay

```

-----
      Percentiles      Smallest
1%                0                0
5%                0                0
10%               0                0      Obs          59707
25%               0                0      Sum of Wgt.  59707

50%               0
                        Largest      Mean          .2056543
75%               0                1      Std. Dev.    .4041823
90%               1                1      Variance     .1633633
95%               1                1      Skewness     1.456512
99%               1                1      Kurtosis     3.121427

```

Pr(lopay=1 assuming u_i=0)

```

-----
      Percentiles      Smallest
1%          .0017309      .0004404
5%          .0041413      .0004943
10%         .0073244      .0005156      Obs          59707
25%         .0196731      .0005494      Sum of Wgt.  59707

50%         .0646826
                        Largest      Mean          .1740193
75%         .221457      .9724768      Std. Dev.    .2350365
90%         .5725967      .9725304      Variance     .0552421
95%         .7588834      .972664      Skewness     1.749767

```

99% .9185546 .9727439 Kurtosis 5.099154

```
.
. *****
. * Do the linear probability model, logit model and probit      *
. * model produce similar predictions of Pr(y=1) for different *
. * types of individuals? Use scatter plots of predicted      *
. * probabilities from the different (random effects) models.  *
. *****
.
. *First plot logit and probit predicted probabilities *
.
. scatter relprob reprob in 1/100, ///
> title(Logit and Probit predicted probabilities) ///
> xtitle(Predicted probabilities from the probit) ///
> ytitle(Predicted probability from the logit) ///
> scheme(slc) legend(on)
.
. graph save "`working'\Logit vs. Probit.gph", replace
(file C:\Users\nw\Documents\HILDA Project\Training\Melbourne Nov
2011\Exercises\Logit vs. Probit.gph saved)
.
. * Then plot LPM and logit predicted probabilities *
. scatter reprob relprob in 1/500, msize(small) ///
> title(LPM and Logit predicted probabilities) ///
> xtitle(Predicted probabilities from the logit) ///
> ytitle(Predicted probabilities from the LPM) scheme(slc)
.
. graph save "`working'\Logit vs. LPM.gph", replace
(file C:\Users\nw\Documents\HILDA Project\Training\Melbourne Nov
2011\Exercises\Logit vs. LPM.gph saved)
.
. *****
. * Estimate the FE logit (recall there is no FE probit).      *
. * Why does Stata report that it has dropped (a lot of!)      *
. * observations? Why is there no constant reported?          *
. *****
. xtlogit lopay hgage agesq cohort married female further degree ///
> jbemp /*tucov*/ permanent nsw, fe
note: multiple positive outcomes within groups encountered.
note: 8873 groups (33576 obs) dropped because of all positive or
all negative outcomes.
note: cohort omitted because of no within-group variance.
note: female omitted because of no within-group variance.

Iteration 0: log likelihood = -9334.0899
Iteration 1: log likelihood = -7311.5943
Iteration 2: log likelihood = -7291.5364
Iteration 3: log likelihood = -7291.4151
Iteration 4: log likelihood = -7291.4151

Conditional fixed-effects logistic regression    Number of obs    =
26131
```

```

Group variable: id
4489

Number of groups =

Obs per group: min =
                avg =
                max =

LR chi2(8) =
5544.70
Log likelihood = -7291.4151
0.0000
Prob > chi2 =

```

```

-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
      hgage |  -0.8804523   0.0268205   -32.83   0.000   -0.9330195   -
0.8278852
      agesq |   0.5477122   0.0333533    16.42   0.000    0.482341
0.6130834
      cohort |   (omitted)
      married |  0.0764314   0.0950317     0.80   0.421   -0.1098272
0.2626901
      female |   (omitted)
      further | -0.3836895   0.113007    -3.40   0.001   -0.6051792   -
0.1621998
      degree | -0.1254846   0.1781733    -0.70   0.481   -0.4746978
0.2237285
      jbempt | -0.0167859   0.0065245    -2.57   0.010   -0.0295737   -
0.0039982
      permanent | 0.1019658   0.0446407     2.28   0.022    0.0144716
0.18946
      nsw |   0.0584974   0.1705381     0.34   0.732   -0.275751
0.3927458
-----
-----

```

```

.
.
. *****
. * Predict the probabilities implied by the FE logit. What do *
. * you notice? How are the probabilities calculated? *
. *****
. predict felprob, pu0

. su felprob, de

```

```

Pr(lopay|fixed effect is 0)
-----
      Percentiles   Smallest
1%      1.85e-15      4.70e-16
5%      5.47e-15      4.90e-16
10%     1.42e-14      5.42e-16
Obs                                59707

```

25%	1.52e-13	5.68e-16	Sum of Wgt.	59707
50%	1.21e-11		Mean	2.08e-07
		Largest	Std. Dev.	8.61e-07
75%	4.77e-09	7.37e-06		
90%	2.34e-07	7.38e-06	Variance	7.42e-13
95%	1.12e-06	7.38e-06	Skewness	5.727186
99%	6.22e-06	7.39e-06	Kurtosis	38.33599

```

.
.
. *****
. * We wish to test the RE assumption that the individual      *
. * effects are uncorrelated with the regressors. We compare  *
. * the RE and FE logit models. Re-estimate both models and   *
. * store the results.                                         *
. *****
. xtlogit lopay hgage agesq cohort married female further degree ///
>      jbemp /*tucov*/ permanent nsw, fe
note: multiple positive outcomes within groups encountered.
note: 8873 groups (33576 obs) dropped because of all positive or
      all negative outcomes.
note: cohort omitted because of no within-group variance.
note: female omitted because of no within-group variance.

```

```

Iteration 0:  log likelihood = -9334.0899
Iteration 1:  log likelihood = -7311.5943
Iteration 2:  log likelihood = -7291.5364
Iteration 3:  log likelihood = -7291.4151
Iteration 4:  log likelihood = -7291.4151

```

```

Conditional fixed-effects logistic regression      Number of obs      =
26131
Group variable: id                               Number of groups   =
4489

Obs per group: min =
2
avg =
5.8
max =
9

LR chi2(8) =
5544.70
Log likelihood = -7291.4151      Prob > chi2      =
0.0000

```

```

-----
-----
lopay |      Coef.   Std. Err.      z    P>|z|     [95% Conf.
Interval]
-----+-----
hgage |  -.8804523   .0268205   -32.83  0.000   -.9330195   -
.8278852
agesq |   .5477122   .0333533    16.42  0.000    .482341
.6130834

```

```

      cohort | (omitted)
    married | .0764314 .0950317 0.80 0.421 -.1098272
.2626901
      female | (omitted)
    further | -.3836895 .113007 -3.40 0.001 -.6051792 -
.1621998
      degree | -.1254846 .1781733 -0.70 0.481 -.4746978
.2237285
      jbempt | -.0167859 .0065245 -2.57 0.010 -.0295737 -
.0039982
      permanent | .1019658 .0446407 2.28 0.022 .0144716
.18946
      nsw | .0584974 .1705381 0.34 0.732 -.275751
.3927458

```

```
-----
-----
```

```
. est store fixed
```

```
.
. xtlogit lopay hgage agesq cohort married female further degree ///
>          jbemp /*tucov*/ permanent nsw, re
```

```
Fitting comparison model:
```

```
Iteration 0: log likelihood = -30339.617
Iteration 1: log likelihood = -23923.118
Iteration 2: log likelihood = -23215.263
Iteration 3: log likelihood = -23201.418
Iteration 4: log likelihood = -23201.391
Iteration 5: log likelihood = -23201.391
```

```
Fitting full model:
```

```
tau = 0.0 log likelihood = -23201.391
tau = 0.1 log likelihood = -22849.103
tau = 0.2 log likelihood = -22527.05
tau = 0.3 log likelihood = -22247.356
tau = 0.4 log likelihood = -22021.286
tau = 0.5 log likelihood = -21862.207
tau = 0.6 log likelihood = -21792.251
tau = 0.7 log likelihood = -21856.716
```

```
Iteration 0: log likelihood = -21792.23
Iteration 1: log likelihood = -21339.586
Iteration 2: log likelihood = -21334.357
Iteration 3: log likelihood = -21334.338
Iteration 4: log likelihood = -21334.338
```

```
Random-effects logistic regression          Number of obs      =
59707
Group variable: id                          Number of groups   =
13362
```

```
Random effects u_i ~ Gaussian              Obs per group: min =
1                                             avg =
```

```
4.5
```

9

max =

5958.40
Log likelihood = -21334.338
0.0000

Wald chi2(10) =

Prob > chi2 =

```
-----
```

	lopay	Coef.	Std. Err.	z	P> z	[95% Conf.	
	Interval]						
	hgage	-.8252997	.0146057	-56.51	0.000	-.8539264	-
.7966731	agesq	.560641	.0145353	38.57	0.000	.5321522	
.5891297	cohort	-.3497448	.0069135	-50.59	0.000	-.3632949	-
.3361947	married	-.0718777	.0500144	-1.44	0.151	-.169904	
.0261487	female	.4493369	.0455548	9.86	0.000	.3600511	
.5386227	further	-.6130987	.049666	-12.34	0.000	-.7104423	-
.515755	degree	-1.300175	.0693114	-18.76	0.000	-1.436023	-
1.164327	jbempt	-.0521376	.0038565	-13.52	0.000	-.0596961	-
.044579	permanent	-.2926862	.0359879	-8.13	0.000	-.3632212	-
.2221512	nsw	-.2433049	.0484902	-5.02	0.000	-.338344	-
.1482658	_cons	708.8378	13.90222	50.99	0.000	681.5899	
736.0856							
	/lnsig2u	1.059181	.0375428			.9855987	
1.132764							
	sigma_u	1.698237	.0318783			1.636892	
1.761881	rho	.4671309	.0093451			.4488672	
.4854832							

```
-----
```

Likelihood-ratio test of rho=0: chibar2(01) = 3734.11 Prob >= chibar2 = 0.000

. est store random

.

* How do the FE coefficients differ from the RE model. Test *
* the FE against the RE model. Which model do you prefer? *

```
. *****
. hausman fixed random
```

----- Coefficients -----				
V_B))	(b)	(B)	(b-B)	sqrt(diag(V_b-
	fixed	random	Difference	S.E.
hgage	-.8804523	-.8252997	-.0551526	.0224947
agesq	.5477122	.560641	-.0129288	.0300194
married	.0764314	-.0718777	.1483091	.0808058
further	-.3836895	-.6130987	.2294092	.101508
degree	-.1254846	-1.300175	1.17469	.1641391
jbmpt	-.0167859	-.0521376	.0353517	.0052628
permanent	.1019658	-.2926862	.394652	.0264133
nsw	.0584974	-.2433049	.3018023	.163499

```
-----
b = consistent under Ho and Ha; obtained from
xtlogit
B = inconsistent under Ha, efficient under Ho; obtained from
xtlogit
```

Test: Ho: difference in coefficients not systematic

$$\begin{aligned} \text{chi2}(8) &= (b-B)' [(V_b - V_B)^{-1}] (b-B) \\ &= 402.83 \\ \text{Prob} > \text{chi2} &= 0.0000 \end{aligned}$$

```
.
.
. *****
. * Predicting the probability for a person with selected *
. * characteristics. *
. * The predict command predicts the probabilities *
. * for each observation separately (given their *
. * characteristics). *
. * To make predictions for a specific type of person we need *
. * to use the model formulae directly and plug in the chosen *
. * values. Use command nlcom (which also gives standard *
. * errors). *
. *
. * Calculate the probability of being low paid for a reference*
. * male aged 40 and born in 1967 (i.e. current yr is 2007), *
. * who is not married, has no further or higher education, *
. * non-unionised permanent job, has been on his job for 1 year*
. * and does not live in nsw.
. * Calculate the probability for
*
. * 3 values of individual effect, low (-1 sd), medium (0) and*
. * high (+1 sd). What impact do unobserved individual level *
. * factors have on the probability of being low paid? *
. *****
.
. * For convenience, use model just estimated (RE logit)
. * "replay" previous results for references
.
```

. xtlogit

Random-effects logistic regression

59707

Group variable: id

13362

Number of obs =

Number of groups =

Random effects u_i ~ Gaussian

1

Obs per group: min =

avg =

4.5

max =

9

Wald chi2(10) =

5958.40

Log likelihood = -21334.338

Prob > chi2 =

0.0000

```
-----  
-----  
          lopay |      Coef.   Std. Err.      z    P>|z|     [95% Conf.  
Interval]-----+-----  
-----  
          hgage |   -.8252997   .0146057   -56.51   0.000   -.8539264   -  
.7966731  
          agesq |    .560641   .0145353    38.57   0.000    .5321522  
.5891297  
          cohort |   -.3497448   .0069135   -50.59   0.000   -.3632949   -  
.3361947  
          married |  -.0718777   .0500144    -1.44   0.151   -.169904  
.0261487  
          female |    .4493369   .0455548     9.86   0.000    .3600511  
.5386227  
          further |  -.6130987   .049666   -12.34   0.000   -.7104423   -  
.515755  
          degree |  -1.300175   .0693114   -18.76   0.000   -1.436023   -  
1.164327  
          jbempt |  -.0521376   .0038565   -13.52   0.000   -.0596961   -  
.044579  
          permanent | -.2926862   .0359879    -8.13   0.000   -.3632212   -  
.2221512  
          nsw |   -.2433049   .0484902    -5.02   0.000   -.338344   -  
.1482658  
          _cons |   708.8378   13.90222    50.99   0.000   681.5899  
736.0856  
-----+-----  
-----  
          /lnsig2u |   1.059181   .0375428                .9855987  
1.132764  
-----+-----  
-----  
          sigma_u |   1.698237   .0318783                1.636892  
1.761881  
          rho |    .4671309   .0093451                .4488672  
.4854832
```

```
-----
-----
Likelihood-ratio test of rho=0: chibar2(01) = 3734.11 Prob >= chibar2 =
0.000
```

```
.
. *****
. * write down xb for ref person using stored coefficients,
. * accessed using _b notation
. * use a local "macro" to store the text for future use
. *****
.
. local refxb _b[_cons] + _b[hgage]*40 + _b[agesq]*16 + ///
> _b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]
.
. * use a macro to store the standard deviation of u(i) *
. local ui=e(sigma_u) //estimate of sd from model
.
. *****
. * NOTE: macro subtlety *
. * A local statement without the = sign stores the *
. * following text directly in the macro *
. * A local statement with the = sign evaluates the expression *
. * and stores the resulting number in the macro *
. * You will see difference below when the macro contents are *
. * displayed. *
. *****
.
. * calculate probabilities using logit formula*
. *low individual effect -use -1 st dev of u(i)*
.
. /*
> nlcom ( prlo:exp(`refxb'-`ui')/(1+exp(`refxb'-`ui')) )
>          ( prmed: exp(`refxb') / (1+exp(`refxb')) )
>          ( prhi: exp(`refxb'+`ui') / (1+exp(`refxb'+`ui')) )
> */
.
. * Note single quotes used to access macro contents
. nlcom prlo:exp(`refxb'-`ui')/(1+exp(`refxb'-`ui'))

          prlo: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]-
1.698237010370481)/ (1+exp(_b[_cons] + _b[hgage]*40
> + _b[agesq]*16 + _b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]-
1.698237010370481))
```

```
-----
-----
```

	lopay	Coef.	Std. Err.	z	P> z	[95% Conf.
Interval]						
	prlo	.0055132	.0003912	14.09	0.000	.0047465
		.0062799				

```
-----
-----
```

```

.
. * medium individual efect, u(o)=0
. nlcom prmed: exp(`refxb') / (1+exp(`refxb'))

      prmed: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]) / (1+exp(_b[_cons] +
_b[hgage]*40 + _b[agesq]*16 +
> _b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]))

```

```

-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
      prmed |   .0294021   .0020361    14.44   0.000   .0254115
.0333928
-----
-----

```

```

.
. * High individual effects - use +1 st dev of u(i)
. nlcom prhi: exp(`refxb'+`ui') / (1+exp(`refxb'+`ui'))

```

```

      prhi: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+1.698237010370481) /
(1+exp(_b[_cons] + _b[hgage]*4
> 0 + _b[agesq]*16 + _b[cohort]*1967 + _b[jbemp]*1 +
_b[permanent]+1.698237010370481))

```

```

-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
      prhi |   .1420207   .0086936    16.34   0.000   .1249814
.1590599
-----
-----

```

```

.
. * calculate xb for ref person who has done further studies *
. local refxb1 _b[_cons] + _b[hgage]*40 + _b[agesq]*16 + ///
> _b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree]

```

```

.
. local ui=e(sigma_u)

```

```

.
. * calculate probabilities *
. nlcom prlo: exp(`refxb1'-`ui') / (1+exp(`refxb1'-`ui'))

```

```

      prlo: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree]-
1.698237010370481) / (1+exp(_b[_cons] +

```

```
> _b[hgage]*40 + _b[agesq]*16 + _b[cohort]*1967 + _b[jbemp]*1 +
_b[permanent]+_b[degree]-1.698237010370481))
```

```
-----
lopay |      Coef.   Std. Err.      z    P>|z|     [95% Conf.
Interval]
-----+-----
prlo |   .0015083   .0001553     9.71   0.000     .001204
.0018126
-----
```

```
. nlcom prmed: exp(`refxb1`) / (1+exp(`refxb1`))
```

```
prmed: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree]) /
(1+exp(_b[_cons] + _b[hgage]*40 + _b[
> agesq]*16 + _b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree]))
```

```
-----
lopay |      Coef.   Std. Err.      z    P>|z|     [95% Conf.
Interval]
-----+-----
prmed |   .0081867   .0008371     9.78   0.000     .0065461
.0098274
-----
```

```
. nlcom prhi: exp(`refxb1'+`ui`) / (1+exp(`refxb1'+`ui`))
```

```
prhi: exp(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 +
_b[permanent]+_b[degree]+1.698237010370481) / (1+exp(_b[_cons] +
> _b[hgage]*40 + _b[agesq]*16 + _b[cohort]*1967 + _b[jbemp]*1 +
_b[permanent]+_b[degree]+1.698237010370481))
```

```
-----
lopay |      Coef.   Std. Err.      z    P>|z|     [95% Conf.
Interval]
-----+-----
prhi |   .0431575   .0042572    10.14   0.000     .0348136
.0515014
-----
```

```
.
*
*****
*
* ADDITIONAL EXERCISE
*
*****
* Repeat probability predictions using RE probit
*
*****
```

```
.
. xtprobit lopay hgage agesq cohort married female further degree ///
>      jbemp /*tucov*/ permanent nsw, re
```

Fitting comparison model:

```
Iteration 0:  log likelihood = -30339.617
Iteration 1:  log likelihood = -23420.818
Iteration 2:  log likelihood = -23251.665
Iteration 3:  log likelihood = -23251.458
Iteration 4:  log likelihood = -23251.458
```

Fitting full model:

```
rho = 0.0      log likelihood = -23251.458
rho = 0.1      log likelihood = -22388.762
rho = 0.2      log likelihood = -21989.801
rho = 0.3      log likelihood = -21836.779
rho = 0.4      log likelihood = -21852.112
```

```
Iteration 0:  log likelihood = -21836.77
Iteration 1:  log likelihood = -21375.42
Iteration 2:  log likelihood = -21367.756
Iteration 3:  log likelihood = -21367.742
Iteration 4:  log likelihood = -21367.742
```

```
Random-effects probit regression          Number of obs      =
59707                                     Number of groups   =
Group variable: id                       Number of groups   =
13362                                     Obs per group: min =
                                           avg =
                                           max =
                                           9
                                           Wald chi2(10)     =
6592.26                                     Prob > chi2       =
Log likelihood = -21367.742
0.0000
```

```
-----+-----
```

	lopay	Coef.	Std. Err.	z	P> z	[95% Conf.	
	Interval]						
	hgage	-.4558991	.0078423	-58.13	0.000	-.4712697	-
.4405285							
	agesq	.3124385	.0079475	39.31	0.000	.2968617	
.3280153							
	cohort	-.1906813	.0037131	-51.35	0.000	-.1979588	-
.1834037							
	married	-.0368732	.0273418	-1.35	0.177	-.0904622	
.0167158							

```
-----+-----
```

```

        female |      .251939   .0251412   10.02   0.000   .2026632
.3012148
        further |     -.3381665   .0274014   -12.34   0.000   -.3918722   -
.2844608
        degree |     -.6975005   .0371214   -18.79   0.000   -.7702571   -
.6247439
        jbempt |     -.0284252   .0020782   -13.68   0.000   -.0324984   -
.024352
        permanent |    -.1737852   .0200036    -8.69   0.000   -.2129916   -
.1345787
        nsw |     -.1308096   .0266997    -4.90   0.000    -.18314   -
.0784791
        _cons |     386.5501    7.463667    51.79   0.000    371.9216
401.1786

```

```

-----+-----
-----
        /lnsig2u |    -.1204002   .0368845                -.1926924   -
.0481079
-----+-----
-----

```

```

        sigma_u |    .9415761   .0173648                .9081496
.976233
        rho |    .4699363   .0091878                .4519754
.4879753
-----+-----
-----

```

Likelihood-ratio test of rho=0: $\chi^2(01) = 3767.43$ Prob $\geq \chi^2 = 0.000$

```

.
. * use a macro to store the standard deviation of u(i) *
. local ui=e(sigma_u) //estimate of sd from model

```

```

.
.
. * calculate probabilities *
. nlcom prlo: norm(`refxb'-'ui')

```

```

        prlo: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]-.9415761176406886)

```

```

-----+-----
-----
        lopay |      Coef.   Std. Err.      z    P>|z|    [95% Conf.
Interval]
-----+-----
        prlo |    .0018619   .0002295     8.11   0.000   .0014122
.0023116
-----+-----
-----

```

```

. nlcom prmed: norm(`refxb')

        prmed: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent])

```

```
-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
-----
      prmed |   .0250518   .0022611    11.08   0.000   .0206202
.0294835
-----
-----
```

. nlcom prhi: norm(`refxb'+`ui')

prhi: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+.9415761176406886)

```
-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
-----
      prhi |   .1544575   .0091814    16.82   0.000   .1364622
.1724527
-----
-----
```

.
.

. * calculate xb for ref person with further studies *

. nlcom prlo: norm(`refxb1'-`ui')

prlo: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree]-
.9415761176406886)

```
-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
-----
      prlo |   .0001602   .0000341     4.70   0.000   .0000934
.0002271
-----
-----
```

. nlcom prmed: norm(`refxb1')

prmed: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 + _b[permanent]+_b[degree])

```
-----
-----
      lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
-----
```

```
prmed | .0039469 .000648 6.09 0.000 .0026769
.005217
```

```
-----
. nlcom prhi: norm(`refxb1'+`ui')
```

```
prhi: norm(_b[_cons] + _b[hgage]*40 + _b[agesq]*16 +
_b[cohort]*1967 + _b[jbemp]*1 +
_b[permanent]+_b[degree]+.9415761176406886)
```

```
-----
lopay |      Coef.   Std. Err.      z    P>|z|      [95% Conf.
Interval]
-----+-----
prhi |      .0431724   .0050745     8.51   0.000   .0332266
.0531183
```

```
-----
.
.
. *****
. *                END OF ADDITIONAL EXERCISE                *
. *****
```

```
. set more on
```

```
. log c
name: <unnamed>
log: C:\Users\nw\Documents\HILDA Project\Training\Melbourne Nov
2011\Exercises\ex6.log
log type: text
closed on: 8 Nov 2011, 09:38:07
```

```
-----
-----
-----
```